Structural and Reduced-form Approaches to Calculating CVA for Portfolios of CDSs

Theory and Practice

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A brief survey of credit literature

Credit Default Swaps (CDSs)

Reduced, structural, and hybrid single- and multi-name models

Credit Value Adjustment (CVA)

Examples
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**Figure:** 5Y CDS Spread in basis points of 5 well known issuers. The spreads are plotted on a different scale before and after 20-Jun-2007. Source: Bank of America - Merrill Lynch Securities.
Figure: ITX and CDX historical On-the-Run Index Spreads. Source: Bank of America - Merrill Lynch Securities.
Figure: On-the-Run ITX Tranche Quote History (0-3% maps to right axis and is % up-front. All other tranche prices are quoted as running spreads and map to the left axis. Since Q1 2009 the 3-6% and 6-9% have also been quoted as % up-front with 500bp running but for the purposes of this chart they have been converted to running spreads for the entire time-series). Prior to Series 9 (20-Mar-2008) the on the run index changes every 6 months, subsequently, index...
CDS modeling

First, we need to explain how to price CDSs.

Then we need to extend our theory to cover indices, tranches, baskets, etc.

Three complementary approaches to pricing CDSs:

reduced form

structural

hybrid

All of them have pros and cons. Reduced form and structural approaches dominate.
We represent intensity of the Cox default process in the form

\[ \tilde{X}(t) = \alpha(t) + \beta X(t) \]

For analytical convenience (rather than for deeper reasons) it is customary to assume that \(X\) is governed by the square-root stochastic differential equation (SDE):

\[ dX(t) = \kappa(\theta(t) - X(t)) \, dt + \sigma \sqrt{X(t)} \, dW(t) + \nu \dd N(t), \quad X(0) = X_0 \]

with exponential (or hyper-exponential) jump distribution.
Survival probability

For practical purposes it is often more convenient to consider discrete jump distributions with jump values $J_m > 0$, $1 \leq m \leq M$, occurring with probabilities $\pi_m > 0$; such distributions are more flexible than parametric ones because they allow one to place jumps where they are needed.

In this framework, the survival probability $q(t, T)$ and the discounted survival probability $Q(t, T)$ from time $t$ to time $T$ has the form

\[
q(t, T) = \mathbb{E}_t \left\{ e^{-\int_t^T \tilde{X}(t') dt'} \right\} = e^{-\int_t^T \alpha(t') dt'} \mathbb{E}_t \left\{ e^{-\int_t^T (\beta X(t')) dt'} \right\}
\]

\[
Q(t, T) = \mathbb{E}_t \left\{ e^{-\int_t^T (r + \tilde{X}(t')) dt'} \right\} = e^{-\int_t^T (r + \alpha(t')) dt'} \mathbb{E}_t \left\{ e^{-\int_t^T (\beta X(t')) dt'} \right\}
\]
Calculation of expectations

Probability $Q(t, T)$ can be computed by solving the following partial differential equation (PDE)

$$(\partial_t + \mathcal{L}) V(t, X, T) - (r + \alpha + \beta X) V(t, X, T) = 0$$

$$V(T, X, T) = 1$$

where

$$\mathcal{L} V \equiv \kappa (\theta - X) V_X + \frac{1}{2} \sigma^2 X V_{XX} + \lambda \left[ \phi \int_0^\infty V(X + J) e^{-\phi J} dJ - V(X) \right]$$

or

$$\mathcal{L} V \equiv \kappa (\theta - X) V_X + \frac{1}{2} \sigma^2 X V_{XX} + \lambda \sum_m \pi_m [V(X + J_m) - V(X)]$$
Affine ansatz 1

The corresponding solution can be written in the so-called affine form:

\[ V(t, X, T) = e^{-\int_{t}^{T} (r + \alpha(t')) dt'} + a(t, T) + b(t, T) X \]

When \( \beta = 1 \), the corresponding \( a, b \) have the form (\( \gamma = \sqrt{\kappa^2 + 2\sigma^2} \))

\[
a = -\frac{2\kappa \theta}{\sigma^2} \ln \left( \frac{\left( (\kappa + \gamma) e^{-(\kappa + \gamma)\tau/2} + (-\kappa + \gamma) e^{-(\kappa - \gamma)\tau/2} \right)}{2\gamma} \right) \\
+ \frac{2\lambda \tau}{(\phi(-\kappa + \gamma) - 2)} \\
- \frac{2\lambda \phi}{(\phi^2 \sigma^2 - 2\phi \kappa - 2)} \ln \left( \frac{(\phi(\kappa + \gamma) + 2) e^{\gamma \tau} + (\phi(-\kappa + \gamma) - 2)}{2\gamma \phi} \right)
\]

\[
b = -\frac{2 (e^{\gamma \tau} - 1)}{(\kappa + \gamma) e^{\gamma \tau} + (-\kappa + \gamma)}
\]
When $\beta \neq 1$, we have to apply scaling

$$
\theta \rightarrow \theta \beta, \quad \sigma \rightarrow \sigma \sqrt{\beta}, \quad \phi \rightarrow \phi / \beta, \quad b \rightarrow b \beta
$$

The survival probability $q(0, T)$, discounted survival probability $Q(0, T)$ and default probability $p(0, T)$ have the form

$$
Q(0, T) = e^{-\int_{t}^{T} (r + \alpha(t')) dt' + a(T) + b(T)X_0}
$$

$$
q(0, T) = e^{-\int_{t}^{T} \alpha(t') dt' + a(T) + b(T)X_0}
$$

$$
p(0, T) = 1 - q(0, T)
$$
The value $U$ of a credit default swap (CDS) paying an up-front amount $\nu$ and a coupon $s$ in exchange for receiving $L = 1 - R$ (where $R$ is the default recovery) on default as follows:

$$U = - \nu + V(0, X_0)$$

Here $V(t, X)$ solves the following pricing problem

$$(\partial_t + \mathcal{L}) V(t, X) - (r + \alpha + \beta X) V(t, X) = s - L(\alpha + \beta X)$$

$$V(T, X) = 0$$

Using Duhamel’s principle and integration by parts, we obtain the following expression for $V$:

$$V(t, X) = -(s + rL) A(t, T) + L[1 - Q(t, T)]$$

where $A(t, T)$ is the annuity factor between two times $t, T$

$$A(t, T) = \int_t^T e^{-\int_t^{t'} (r+\alpha(t'')) dt''} + a(t, t') + b(t, t') X dt' = \int_t^T Q(t, t') dt$$
CDS valuation 2

For a given up-front payment \( v \), we can represent the corresponding par spread \( \hat{s} \) (i.e. the spread which makes the value of the corresponding CDS zero) as follows:

\[
\hat{s}(T) = \frac{-v + L[1 - Q(0, T) - rA(0, T)]}{A(0, T)}
\]

Conversely, for a given spread we can represent the par up-front payment in the form

\[
\hat{v} = -(s + rL)A(0, T) + L[1 - Q(0, T)]
\]

In these formulas we implicitly assumed that the corresponding CDS is fully collateralized, so that in the event of default \( L \) is readily available.

We can use these formulas to calibrate \( \alpha(t) \) to a given spread curve \( s(t) \).
Two name correlation

It is very tempting to extend the above framework to cover several correlated names. For example, consider two credits, 1, 2 and assume for simplicity that their default intensities have the form,

\[ X_i(t) = \alpha_i(t) + \beta_i X(t) \]

and their losses given default are \( L_i \).

For a given maturity \( T \), the default event correlation \( \rho \) is defined as follows

\[
\rho(0, T) = \frac{p_{12}(0, T) - p_1(0, T)p_2(0, T)}{\sqrt{p_1(0, T)(1 - p_1(0, T))}p_2(0, T)(1 - p_2(0, T))}
\]

where \( \tau_1, \tau_2 \) are the default times, and

\[
p_1(0, T) = P(\tau_1 \leq T), \quad p_1(0, T) = P(\tau_2 \leq T)
\]

\[
p_{12}(0, T) = P(\tau_1 \leq T, \tau_2 \leq T)
\]
It is clear that
\[ p_i(0, T) = 1 - e^{-\int_0^T \alpha_i(t') dt'} + a(0, T; \beta_i) + b(0, T; \beta_i) X_0 \]

Simple calculation yields
\[ p_{12}(0, T) = 1 - q_1(0, T) - q_2(0, T) + q_{12}(0, T) \]

where
\[ q_{12}(0, T) = e^{-\int_0^T (\alpha_1(t') + \alpha_2(t')) dt'} + a(0, T; \beta_1 + \beta_2) + b(0, T; \beta_1 + \beta_2) X_0 \]

Thus
\[ \rho(0, T) = \frac{q_{12}(0, T) - q_1(0, T) q_2(0, T)}{\sqrt{p_1(0, T) q_1(0, T) p_2(0, T) q_2(0, T)}} \]

In the absence of jumps, the corresponding event correlation is very low. If large positive jumps are added (while overall survival probability is preserved), then correlation can increase all the way to one. We illustrate this observation below.
Figure: Correlation $\rho$ and mean-reversion level $\theta = X_0$ as functions of jump intensity $\lambda$. Other parameters are as follows: $T = 5y$, $\kappa = 0.5$, $\sigma = 7\%$, and $J = 5.0$. 
In the two-name portfolio, we can define two types of CDSs which depend on the correlation: (A) the first-to-default swap (FTD); (B) the second-to-default swap (STD). The corresponding par spreads $\hat{s}_1, \hat{s}_2$ can be found by using discounted survival probabilities of the form (we are assuming that $L_1 = L_2 = L$)

\[
Q_{12}(0, T) = e^{-\int_0^T (r + \alpha_1(t') + \alpha_2(t')) dt'} + a(0, T; \beta_1 + \beta_2) + b(0, T; \beta_1 + \beta_2) X_0
\]

\[
\hat{Q}_{12}(0, T) = Q_1(0, T) + Q_2(0, T) - Q_{12}(0, T)
\]

It is clear that the relative values of $\hat{s}_1, \hat{s}_2$ very strongly depend on whether or not jumps are present in the model.
Figure: FTD spread $\hat{s}_1$, STD spread $\hat{s}_2$, and single name CDS spread $\hat{s}$ as functions of jump intensity $\lambda$. Other parameters are the same as in Fig.1. It is clear that jumps are necessary to have $\hat{s}_1$ and $\hat{s}_2$ of similar magnitudes.
Counter-party effects

An important application of the above model is to the evaluation of counter-party effects on fair CDS spreads. Consider three names, 1, 2, 3, and assume that name 2 has written a CDS on reference name 3, which is bought by name 1. For simplicity, we assume that their spreads have the form

\[ X_i = \alpha_i + \beta_i X \]

It is clear that the pricing problem for the value of the uncollateralized CDS \( V^{\{2,3\}} \) can be written as follows

\[
\mathcal{L} V^{\{2,3\}} (t, X) - \left( r + \alpha_2 + \alpha_3 + (\beta_2 + \beta_3) X \right) V^{\{2,3\}} (t, X) \\
= s - L_3 (\alpha_3 + \beta_3 X) - \left( R_2 U^{\{3\}}_+ (t, X) + U^{\{3\}}_- (t, X) \right) (\alpha_2 + \beta_2 X)
\]

where \( U^{\{3\}} \) is the value of a fully collateralized CDS on name 3. It is clear that the discount rate is increased from \( r + \alpha_3 + \beta_3 X \) to \( r + \alpha_2 + \alpha_3 + (\beta_2 + \beta_3) X \), since there are two cases when the uncollateralized CDS can be terminated due to default: when the reference name 3 defaults; when the issuer 2 defaults.
Alternatively, the pricing equation can be written in terms of CVA_1 as follows

\[ \mathcal{L} W^{2,3} (t, X) - (r + \alpha_2 + \alpha_3 + (\beta_2 + \beta_3) X) W^{2,3} (t, X) \]

\[ = L_2 U_{+-}^{3} (t, X) (\alpha_2 + \beta_2 X) \]

where \( W^{2,3} = V^{2,3} - U^{3} \). It is clear that in this formulation in addition to coupon payments at rate \( s \); the buyer 1 has to account for additional losses incurred when the issuer 2 defaults.

Although this equation is no longer analytically solvable, it can be solved numerically via, say, an appropriate modification of the classical Crank-Nicholson method. It turns out that in the presence of jumps the value of the fair par spread goes down dramatically.
If the buyer is risky, the pricing problem for $V^{1,3}$ can be written as follows

$$
\mathcal{L} \mathcal{W}^{1,3} (t, X) - (r + \alpha_1 + \alpha_3 + (\beta_1 + \beta_3) X) \mathcal{W}^{1,3} (t, X)
= L_1 U_3^- (t, X) (\alpha_1 + \beta_1 X)
$$

If all three names are risky, the pricing problem is symmetric and quite intuitive

$$
\mathcal{L} \mathcal{W}^{1,2,3} (t, X) - (r + \alpha_1 + \alpha_2 + \alpha_3 + (\beta_1 + \beta_2 + \beta_3) X) \mathcal{W}^{1,2,3} (t, X)
= L_1 U_3^- (t, X) (\alpha_1 + \beta_1 X) + L_2 U_3^+ (t, X) (\alpha_2 + \beta_2 X)
$$
Figure: $r = 3\%, s_1 = 3\%, R_1 = 40\%, \beta_1 = 1, x_1 = 3\%, s_2 = 3\%, R_2 = 40\%, \beta_2 = 1.5, x_2 = 3\%, s_3 = 3\%, R_3 = 40\%, \beta_3 = 1.6, x_3 = 2.5\%, \kappa = 0.5, \theta = 2\%, \sigma = 10\%, \lambda = 2\%, \phi = 1, T = 5Y$
In case of a *portfolio* of $N - 2$ CDSs the pricing problem becomes extraordinary complex since $2^{N-2} - 1$ pricing problems need to be solved. In practice, with proper parallelization we can handle $N \sim 20$ via PDEs. For larger $N$ we have to use MC. Let’s assume that $N = 4$, i.e. we deal with two CDSs. Then

$$\mathcal{L} W^{\{1,2,3,4\}} (t, X) - (r + X_1 + X_2 + X_3 + X_4) W^{\{1,2,3,4\}} (t, X)$$

$$= L_1 X_1 U^{\{3,4\}}_-(t, X) + L_2 X_2 U^{\{3,4\}}_+(t, X) - X_3 W^{\{1,2,4\}} - X_4 W^{\{1,2,3\}}$$

Thus, in order to compute $W^{\{1,2,3,4\}}$ we need to compute $W^{\{1,2,4\}}, W^{\{1,2,3\}}$ first.
We can split the problem as follows. Define the following combinations

\[ R_{\pm}^{3,4} = U_{\pm}^{3,4} - U_{\pm}^3 - U_{\pm}^4 \]

\[ S_{\{1,2,3,4\}} = W_{\{1,2,3,4\}} - W_{\{1,2,3\}} - W_{\{1,2,4\}} \]

we obtain the following system of decoupled equations which emphasizes portfolio effects

\[ \mathcal{L} S_{\{1,2,3,4\}} (t, X) - (r + X_1 + X_2 + X_3 + X_4) S_{\{1,2,3,4\}} (t, X) \]

\[ = L_1 X_1 R_{-}^{3,4} (t, X) + L_2 X_2 R_{+}^{3,4} (t, X) \]

Finally

\[ W_{\{1,2,3,4\}} = S_{\{1,2,3,4\}} + S_{\{1,2,3\}} + S_{\{1,2,4\}} \]
Traditionally, square-root process causes a lot of confusion when boundary conditions are discussed. However, recently it was clarified (see, e.g., Ekstrom et al. 2009)
The so-called Fichera condition suggest that no boundary condition is needed when

\[
\lim_{X \to 0} \left( \text{drift} (x) - \frac{1}{2} \frac{d}{dx} (\text{variance} (x)) \right) \geq 0
\]

For square-root process this is equivalent to the Feller condition. Still, modern practice suggest that using equation as boundary condition always works:

\[
V_t (t, 0) + \kappa (\theta - X) V_X (t, 0) = 0
\]

We also use equation as a far-field condition.
We derive the usual FD discretization which is almost tri-diagonal but the left and right blocks are $3\times3$ and $4\times4$ matrices, respectively, which can easily be handled.

We treat the jump term fully explicitly as in Lipton (2003). Specifically

\[
I(x) = e^{-\phi dx} I(x + dx) + \frac{e^{-\phi dx} - 1 + \phi dx}{\phi dx} V(x) + \frac{1 - (1 + \phi dx) e^{-\phi dx}}{\phi dx} V(x)
\]

The corresponding scheme is very fast and robust.
Consider a company and assume that its asset value is driven by the following SDE

\[
\frac{dV(t)}{V(t)} = \left( r - d - \kappa \lambda(t) \right) dt + \sigma(t) dW(t) + \left( e^J - 1 \right) dN(t) \quad V_i(0) = v_i
\]

\[
\kappa = \mathbb{E} \left\{ e^J - 1 \right\}
\]

Default boundary is driven by the following ODE

\[
\frac{dB(t)}{B(t)} = \left( r - d \right) dt \quad B(0) = b
\]

In log coordinates

\[
x(t) = \ln \left( \frac{V(t)}{B(t)} \right)
\]

\[
dx(t) = \mu(t) dt + \sigma(t) dW(t) + J dN(t) \quad x(0) = \ln \left( \frac{v}{b} \right) = \xi
\]

\[
\mu = -\frac{1}{2} \sigma^2 - \kappa \lambda
\]

The corresponding default boundary is flat
Thus, the asset value is governed by a combination of a Wiener process and a Poisson process with exponentially distributed jumps. The firm defaults if the value $x(t)$ crosses zero barrier $b(t) = 0$. It was realized early on that without jumps (or/and curvilinear or uncertain barriers) it is impossible to explain the short end of the CDS curve within the structural framework. When all the relevant parameters are constant, the problem can be solved analytically via the Laplace transform. However, in general this approach does not work.
Unbounded PDF figure

\[ \text{sigma}=0.2, \lambda=0.1, \alpha=1.0, \mu=0.0 \]
Hitting time density figure

Hitting time density

Time
Prob
Jump-Diffusion Inversion
Regular Diffusion Inversion
Regular Diffusion Analytical

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CVA for CDSs
Perturbative calculation of Green’s function in 1D. For brevity we assume that all the parameters are time independent. Time-dependent case can be solved in a similar way. To start with, we make the following transform

\[ \Phi(t, x) = \exp \left( - \left( \frac{\mu^2}{2\sigma^2} + \lambda \right) t + \frac{\mu}{\sigma^2} (x - \xi) \right) \tilde{\Phi}(t, x) \]

We assume that \( \lambda \ll 1 \) and represent \( \tilde{\Phi} \) as follows

\[ \tilde{\Phi}(t, x) = \tilde{\Phi}^{(0)}(t, x) + \lambda \tilde{\Phi}^{(1)}(t, x) + \ldots \]

A simple calculation yields (with \( \tau = \sigma^2 t \))

\[ \tilde{\Phi}^{(0)}(t, x) = \frac{1}{\sqrt{\tau}} \left[ n \left( \frac{x - \xi}{\sqrt{\tau}} \right) - n \left( \frac{x + \xi}{\sqrt{\tau}} \right) \right] \]
A very elaborate calculation yields

\[ \tilde{\Phi}^{(1)} (t, x) = \frac{\alpha}{\bar{\alpha}\sigma_1^2} \left[ \bar{\alpha}\tau P \left( -\frac{x - \xi}{\sqrt{\tau}}, -\bar{\alpha}\sqrt{\tau} \right) \right. \]

\[ + xP \left( -\frac{x + \xi}{\sqrt{\tau}}, -\bar{\alpha}\sqrt{\tau} \right) - (x - \bar{\alpha}\tau) P \left( -\frac{x + \xi}{\sqrt{\tau}}, \bar{\alpha}\sqrt{\tau} \right) \]

\[ - (x + \bar{\alpha}\tau) e^{-\bar{\alpha}\xi} P \left( -\frac{x}{\sqrt{\tau}}, -\bar{\alpha}\sqrt{\tau} \right) + (x - \bar{\alpha}\tau) e^{-\bar{\alpha}\xi} P \left( -\frac{x}{\sqrt{\tau}}, \bar{\alpha}\sqrt{\tau} \right) \]

The formula can be independently checked via the method of heat potentials.
Figure: $G(t, x)$ for $t = 20$, $\sigma = 20\%$, $\lambda = 2\%$, $\alpha = 4$, $\zeta = 1$. 

Quality of approximation figure
Barrier PDF figure

\[ \text{sigma}=0.2, \text{lambda}=0.1, \text{alpha}=1.0, \text{mu}=0.0, \text{barrier}=-2.0 \]

\[ 0.00 \quad 0.10 \quad 0.20 \quad 0.30 \quad 0.40 \quad 0.50 \quad 0.60 \quad 0.70 \]

\[ -5.00 \quad -4.00 \quad -3.00 \quad -2.00 \quad -1.00 \quad 0.00 \quad 1.00 \quad 2.00 \quad 3.00 \quad 4.00 \quad 5.00 \]

\[ X \]

\[ \text{Prob} \]

\[ \text{Series PDE Laplace} \]

\text{A Lipton (Bank of America Merrill Lynch) CVA for CDSs 29/10 39 / 61}
Comparison of numerical and analytical t.p.d.f. as well as the solution of the barrier problem with non-constant barrier is given below.

By bootstrapping the barrier, we can reproduce term structure of CDS for most names. In addition, we can price equity derivatives and produce a respectable volatility skew.

The barrier can assumed to be stochastic as in Duffie and Lando (2001) or CreditGrades (2000). Reduction of the information set makes reduced form and structural modeling almost identical, Jarrow, Protter, Yildirim (2004).
We need to analyze behavior of two correlated Brownian motions in a quarter-plane.

Alternatively, we can map a quarter-plane into an angle, \( \alpha = \text{arctan} (-\tilde{\rho} / \rho) \).

There are two methods to solve the problem: (A) method of eigenfunction expansion; (B) method of images. In case (A) we have

\[
G_{\alpha}(t, r, \phi | 0, r', \phi') = \frac{2e^{-(r^2 + r'^2) / 2t}}{\alpha t} \sum_{n=1}^{\infty} I_{n\pi / \alpha} \left( \frac{rr'}{t} \right) \sin \left( \frac{n\pi \phi'}{\alpha} \right) \sin \left( \frac{n\pi \phi}{\alpha} \right)
\]

In case (B) we introduce the following function \( f(p, q) \) with \( p \geq 0, -\infty < q < \infty \), (which is a close relation of \( \alpha \)-stable Levy distributions):

\[
f(p, q) = \frac{1}{\sqrt{2\pi}} \left( \frac{1}{2} e^{-p(cos(q) + 1)} - \frac{1}{2\pi} \int_{0}^{\infty} e^{-p(cosh(q\zeta) + 1)} \frac{d\zeta}{\zeta^2 + 1} \right)
\]
We also define the following "derivative" function $h(p, \phi)$:

$$h(p, \phi) = s_+ f(p, \pi + \phi) + s_- f(p, \pi - \phi)$$

Then we can represent unbounded $G(t, r, \phi|0, r', \phi')$ in the following remarkable form, which can be viewed as a direct generalization of the planar 2D case:

$$G(t, r, \phi|0, r', \phi') = \frac{1}{\sqrt{t}} g \left( \frac{r - r'}{\sqrt{t}} \right) \frac{1}{\sqrt{t}} h \left( \frac{rr'}{t}, \phi - \phi' \right)$$

The corresponding $G_\alpha$ can be written as follows:

$$G_\alpha(t, r, \phi|0, r', \phi') = \sum_{n=-\infty}^{\infty} G(t, r, \phi|0, r', \phi' + 2n\alpha)$$

$$- \sum_{n=-\infty}^{\infty} G(t, r, \phi|0, r', -\phi' + 2n\alpha)$$
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Three- and Multi-name case

Consider \( N \) companies and assume that their asset values are driven by the following SDEs

\[
\frac{dV_i(t)}{V_i(t)} = (r - d_i - \kappa_i \lambda_i(t)) \, dt + \sigma_i(t) \, dW_i(t) + \left(e^{J_i} - 1\right) \, dN_i(t) \quad V_i(0)
\]

\[
\kappa_i = \mathbb{E} \left\{ e^{J_i} - 1 \right\}
\]

Default boundaries are driven by the ODEs of the form

\[
\frac{dB_i(t)}{B_i(t)} = (r - d_i) \, dt \quad B_i(0) = b_i
\]

In log coordinates

\[
x_i(t) = \ln \left( \frac{V_i(t)}{B_i(t)} \right)
\]

\[
dx_i(t) = \mu_i(t) \, dt + \sigma_i(t) \, dW_i(t) + J_i \, dN_i(t) \quad x_i(0) = \ln \left( \frac{v_i}{b_i} \right) = \zeta_i
\]

\[
\mu_i = -\frac{1}{2} \sigma_i^2 - \kappa_i \lambda_i
\]
Correlation structure

We introduce correlation between diffusions in the usual way and assume that

\[ dW_i(t) \, dW_j(t) = \rho_{ij}(t) \, dt \]

We introduce correlation between jumps following the Marshall-Olkin idea. Let \( \Pi^{(N)} \) be the set of all subsets of \( N \) names except for the empty subset \( \{\emptyset\} \), and \( \pi \) its typical member. With every \( \pi \) we associate a Poisson process \( N_\pi(t) \) with intensity \( \lambda_\pi(t) \), and represent \( N_i(t) \) as follows

\[
N_i(t) = \sum_{\pi \in \Pi^{(N)}} 1\{i \in \pi\} N_\pi(t)
\]

\[
\lambda_i(t) = \sum_{\pi \in \Pi^{(N)}} 1\{i \in \pi\} \lambda_\pi(t)
\]

Thus, we assume that there are both common and idiosyncratic jump sources.
Multi-name pricing problem

We now formulate a typical pricing equation in the positive cone $R_+^{(N)}$. We have

$$\partial_t U (t, \vec{x}) + L^{(N)} U (t, \vec{x}) = \chi (t, \vec{x})$$

$$U (t, \vec{x}_{0,k}) = \phi_{0,k} (t, \vec{y}), \quad U (t, \vec{x}_{\infty,k}) = \phi_{\infty,k} (t, \vec{y})$$

$$U (T, \vec{x}) = \psi (\vec{x})$$

where $\vec{x}_{0,k}, \vec{x}_{\infty,k}, \vec{y}_k$ are $N$ and $N - 1$ dimensional vectors, respectively,

$$\vec{x}_{0,k} = \left( x_1, ..., 0, ... x_N \right), \quad \vec{x}_{\infty,k} = \left( x_1, ..., \infty, ... x_N \right), \quad \vec{y}_k = (x_1, ..., x_N)$$

The integro-differential operator $L^{(N)}$ can be written as

$$L^{(N)} f (\vec{x}) = \frac{1}{2} \sum_i \sigma_i^2 \partial_i^2 f (\vec{x}) + \sum_{i,j,j > i} a_{ij} \partial_i \partial_j f (\vec{x}) + \sum_i \beta_i \partial_i f (\vec{x})$$

$$- \tilde{\gamma} f (\vec{x}) + \sum_{\pi \in \Pi^{(N)}} \lambda_{\pi} \prod_{i \in \pi} J_i f (\vec{x})$$

Here $a_{ij} = \sigma_i \sigma_j \rho_{ij}$, $\tilde{\gamma} = \gamma + \sum_{\pi \in \Pi^{(N)}} \lambda_{\pi}$. 
Jump terms

In the case of negative exponential jumps,

\[ J_i f (\bar{x}) = \phi_i \int_{-x_i}^{0} f (x_1, ..., x_i + j, ..., x_N) e^{-\phi_i j} dj \]

while in the case of discrete negative jumps

\[ J_i f (\bar{x}) = H (x_i + J_i) f (x_1, ..., x_i + J_i, ..., x_N) \]

and \( H \) is the Heaviside function.

We naturally split the operator \( \mathcal{L}^{(N)} \) into the local (differential) and non-local (integral) parts:

\[ \mathcal{L}^{(N)} f (\bar{x}) = \mathcal{D}^{(N)} f (\bar{x}) + \mathcal{I}^{(N)} f (\bar{x}) \]
The corresponding adjoint operator is

\[
\mathcal{L}^{(N)\dagger} g(\tilde{x}) = \frac{1}{2} \sum_i \sigma^2_i \partial^2_i g(\tilde{x}) + \sum_{i,j,j > i} a_{ij} \partial_i \partial_j g(\tilde{x}) - \sum_i \beta_i \partial_i g(\tilde{x}) - \tilde{\gamma} g(\tilde{x}) + \sum_{\pi \in \Pi^{(N)}} \lambda^\pi \prod_{i \in \pi} J_i^\dagger g(\tilde{x})
\]

where

\[
J_i^\dagger g(\tilde{x}) = \int_{0}^{\infty} g(x_1, \ldots, x_i - j, \ldots x_N) \omega_i(j) \, dj
\]

\[
J_i^\dagger g(\tilde{x}) = g(x_1, \ldots, x_i - J_i, \ldots x_N)
\]
Green’s formula

We introduce Green’s function $\Phi(t', \bar{x})$, or, more explicitly, $\Phi(t', \bar{x}; t, \bar{\zeta})$, such that

$$\partial_{t'} \Phi(t', \bar{x}) - \mathcal{L}^{(N)} \Phi(t', \bar{x}) = 0$$

$$\Phi(t', \bar{x}_{0k}) = 0, \quad \Phi(t', \bar{x}_{\infty k}) = 0, \quad \Phi(t, \bar{x}) = \delta(\bar{x} - \bar{\zeta})$$

It can be shown that

$$U(t, \bar{\zeta}) = \int_{\mathcal{R}_+^{(N)}} \psi(\bar{x}) \Phi(T, \bar{x}; t, \bar{\zeta}) \, d\bar{x}$$

$$+ \sum_{k} \int_{t}^{T} \int_{\mathcal{R}_+^{(N-1)}} \phi_k(t', \bar{y}) \left\{ \frac{1}{2} \sigma_k^2 \Phi_k(t', \bar{y}; t, \bar{\zeta}) \right\} \, dt' \, d\bar{y}$$

$$- \int_{t}^{T} \int_{\mathcal{R}_+^{(N)}} \chi(t', \bar{x}) \Phi(t', \bar{x}; t, \bar{\zeta}) \, dt' \, d\bar{x}$$

In other words, instead of solving the backward pricing problem with nonhomogeneous rhs and boundary conditions, we can solve the forward propagation problem for Green’s function with homogeneous rhs and boundary conditions.
Non-risky CDS pricing problem

We are now in a position to formulate the pricing problems of interest.

We start with a CDS on a reference credit 2 which is sold by a nonrisky seller 3 to a nonrisky buyer 1. Assuming for simplicity that the coupon is paid continuously at the rate $s$, and that the default boundary is continuous, we write the pricing equation as follows

$$
\partial_t U^{(2)}(t, x_2) + \mathcal{L}^{(2)} U^{(2)}(t, x_2) = \chi(t, x_2)
$$

$$
U^{(2)}(t, 0) = \phi_0(t), \quad U^{(2)}(t, \infty) = \phi_\infty(t)
$$

$$
U^{(2)}(T, x_2) = 0
$$

Here

$$
\chi(t, x_2) = s - \lambda_2 (1 - \hat{R}_2) e^{-\alpha_2 x_2} \quad \hat{R}_2 = \frac{\alpha_2 R_2}{\alpha_2 + 1}
$$

$$
\phi_0(t) = (1 - R_2), \quad \phi_\infty(t) = -s \int_t^T D(t, t') \, dt'
$$
For a CDS on a reference credit 2 which is sold by a risky seller 3 to a nonrisky buyer 1, we write the pricing equation as follows

\[ \partial_t U^{(2,3)}(t, x_2, x_3) + \mathcal{L}^{(2,3)} U^{(2,3)}(t, x_2, x_3) = \chi(t, x_2, x_3) \]

\[ U^{(2,3)}(t, 0, x_3) = \phi_{0,2}(t, x_3), \quad U^{(2,3)}(t, \infty, x_3) = \phi_{\infty,2}(t, x_3) \]

\[ U^{(2,3)}(t, x_2, 0) = \phi_{0,3}(t, x_2), \quad U^{(2,3)}(t, x_2, \infty) = \phi_{\infty,3}(t, x_2) \]

\[ U^{(2,3)}(T, x_2, x_3) = 0 \]
## Table 1. Market data, as of October 2, 2009, for JPM and MS.

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<th></th>
<th>( s(0) )</th>
<th>( L(0) )</th>
<th>( R )</th>
<th>( I(0) )</th>
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<td>41.77</td>
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<td>203</td>
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<td>DF</td>
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<td>1y</td>
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<td>1.10</td>
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Table 2. CDS spreads and other relevant outputs for JPM and MS.
Figure: Calibrated intensity rates for JMP (lhs) and MS(rhs) in the models with ENJs and DNJs, respectively.
Figure: PDF of the driver $x(T)$ for JMP in the model with DNJs (lhs) and ENJs (rhs) for $T = 1y, 5y, \text{ and } 10y$. 
Figure: Same graphs as in Figure 3 but for MS.
Figure: Log-normal credit default swaption volatility implied from model with $T_e = 1y$, $T_t = 5y$ as a function of the inverse moneyness $K / \bar{c}(T_e, T_e + T_t)$. 
Figure: Log-normal equity volatility implied by the model as a function of inverse moneyness $K/s$ for put options with $T = 0.5y$. 
Figure: Equilibrium spread for protection buyer and protection seller of a CDS on MS with JPM as the counterparty.
Figure: Same for a CDS on JPM with MS as the counterparty.
We gave a broad overview of the state of CDS modeling.

We showed how to build meaningful reduced-form and structural default models and connected them to each other.

We showed that even the simplest elementary building blocks of credit universe are difficult to describe properly and require a lot of effort.

The opinions expressed in this talk are those of the speaker and do not necessarily reflect the views or opinions of Bank of America Merrill Lynch.